

# Lecture 10

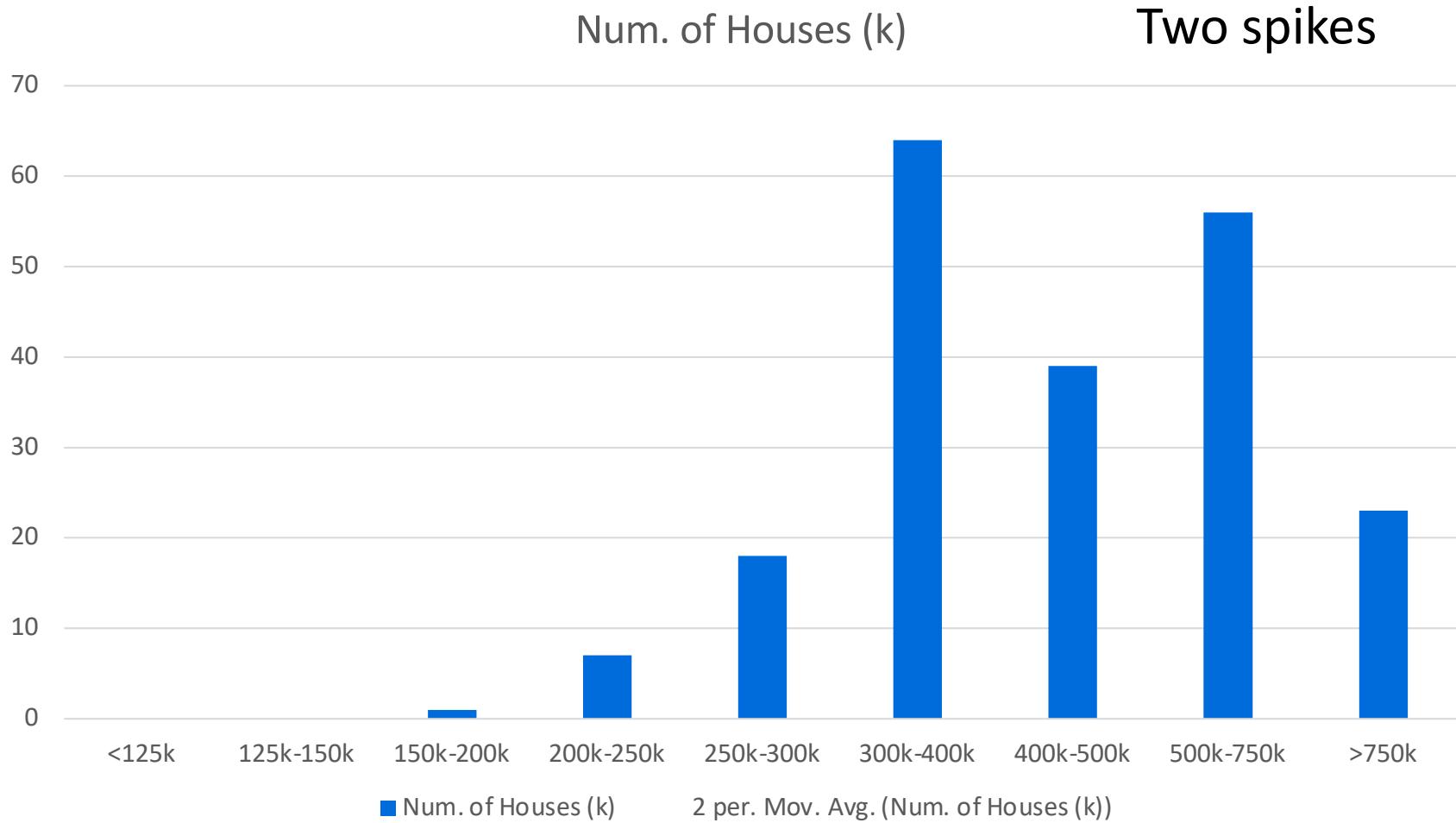
# Gaussian Mixture Models

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# Recap

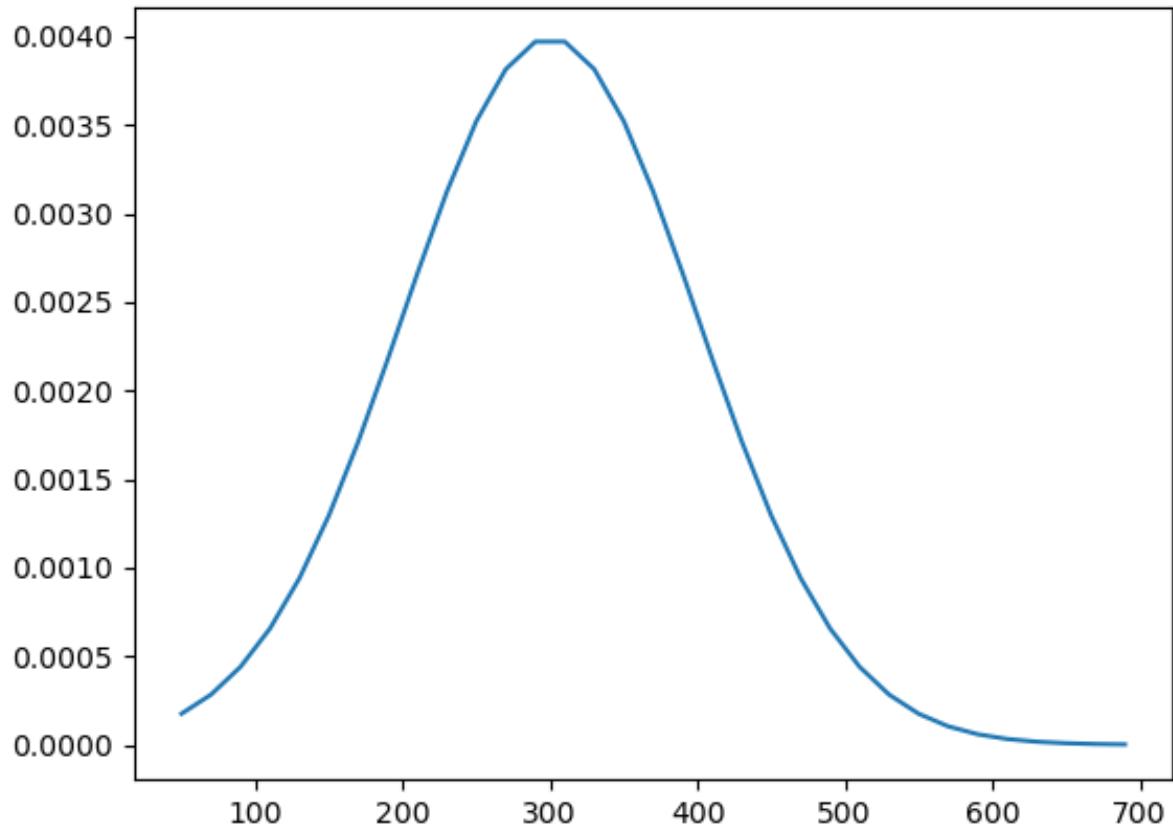
- Bayesian networks:
  - Directed acyclic graph
  - Nodes are random variables
  - arcs are probabilistic dependencies
- Examine dependence of two variables given observation: d-separation

# Housing Price Pattern



# Gaussian Distribution

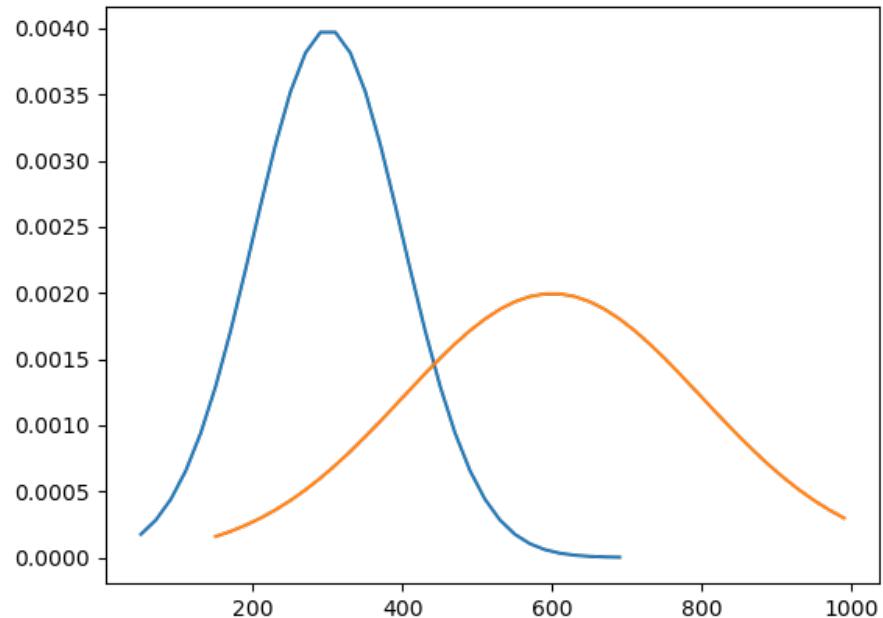
Only single spike



$$p(x) = \frac{1}{\sqrt{(2\pi)^k \det(\Sigma)}} \exp\left[-\frac{1}{2}(x - \mu)^T \Sigma^{-1} (x - \mu)\right]$$

# Two Underlying Patterns

- It might be multiple underlying patterns of Gaussian distribution
  - Los Angeles and Pittsburgh have different median housing price



# Gaussian Mixture Model

Generative process:

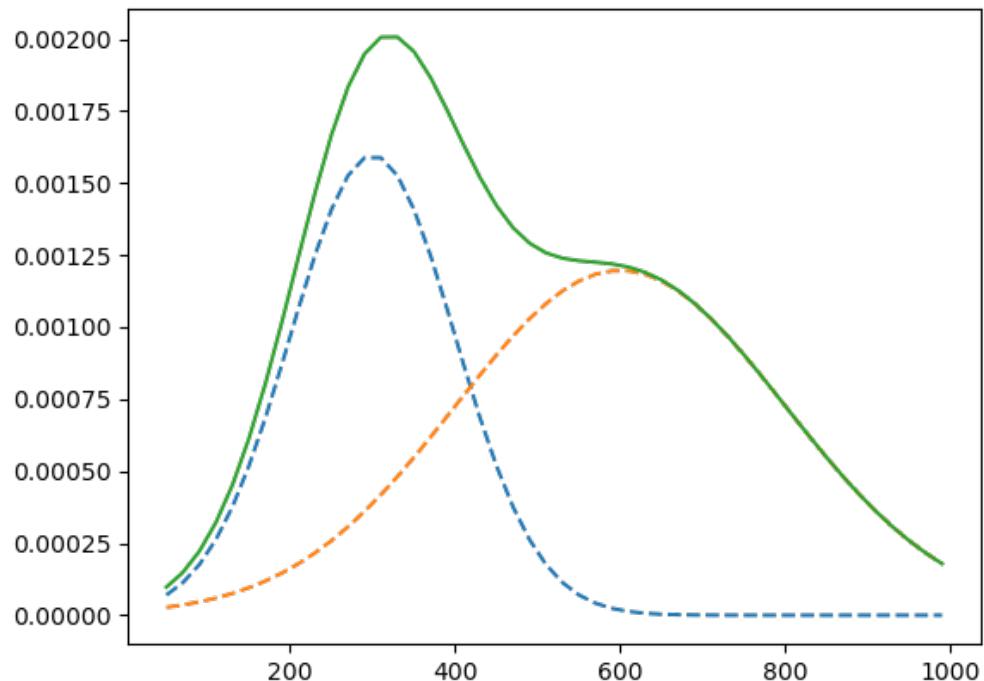
- $z \sim \text{Categorical}(K)$
- $x|z \sim \text{Gaussian}(\mu_z, \Sigma_z)$
- Density:

$$\begin{aligned} p(z, x) &= p(z) \cdot p(x|z) \\ &= \left\{ \begin{array}{l} w_0 \cdot N(x|\mu_0, \Sigma_0) \\ w_1 \cdot N(x|\mu_1, \Sigma_1) \end{array} \right. \end{aligned}$$

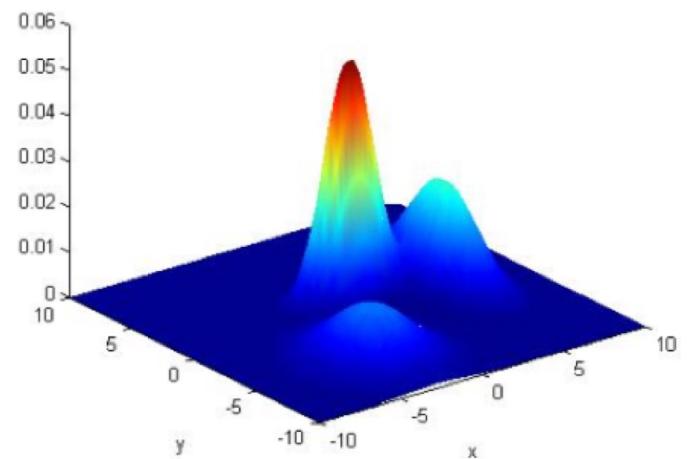
$\sum_{i=1}^K p(z=i, x) = \sum_{i=1}^K p(z=i)p(x|z=i)$

$$= w_0 N(x|\mu_0, \Sigma_0) + w_1 N(x|\mu_1, \Sigma_1)$$

# Gaussian Mixture



**d=2:**



# Mixture Distribution

- Z: latent variable
- $x|z$  can be any distribution in parametric form (e.g. exponential distribution)



# Learning Parameters for GMM

- Observation:  $x_{1..N}$
- $\theta = \{w_{1..k}, \mu_{1..k}, \Sigma_{1..k}\}$
- MLE (with latent variable z)
- Log-likelihood:
- Expectation-maximization algorithm

$$\begin{aligned} p(z, x) &= p(z) \cdot p(x|z) \\ &= \left\{ \begin{array}{l} w_0 \cdot N(x|\mu_0, \Sigma_0) \\ w_1 \cdot N(x|\mu_1, \Sigma_1) \end{array} \right. \end{aligned}$$

(z)  
X

$$\begin{aligned} p(x) &= \sum_{i=1}^K p(z=i, x) = \sum_{i=1}^K p(z=i)p(x|z_i) \\ &= w_0 N(x|\mu_0, \Sigma_0) + w_1 N(x|\mu_1, \Sigma_1) \end{aligned}$$

Optimality condition

$$\begin{aligned} \mathcal{L}(\theta) &= \log \prod_{n=1}^N p(x_n | \theta) \\ &= \sum_{n=1}^N \log \sum_{i=1}^K p(z_n=i) \cdot p(x_n; \mu_i, \Sigma_i) \end{aligned}$$

taking  $\frac{\partial \mathcal{L}(\theta)}{\partial \theta} = 0$  no closed form solution

# Expected log-likelihood

- $L(\theta) = E_{p(z_n|x_n)} \log p(x_n, z_n)$

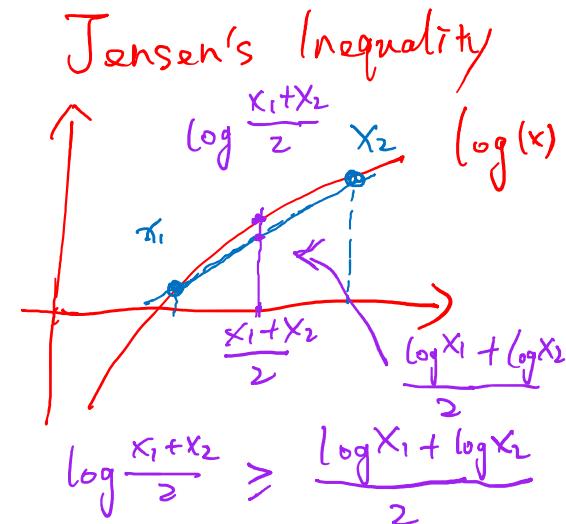
$$\ell(\theta) = \sum_{n=1}^N \log \sum_{i=1}^K p(z_n=i) \cdot p(x_n | z_n=i)$$

$$= \sum_{n=1}^N \log \sum_{i=1}^K p(z_n=i | x_n) \cdot \frac{p(z_n=i) \cdot p(x_n | z_n=i)}{p(z_n=i | x_n)}$$

Jensen

$$\geq \sum_{n=1}^N \sum_{i=1}^K p(z_n=i | x_n) \cdot \log \frac{p(z_n=i) \cdot p(x_n | z_n=i)}{p(z_n=i | x_n) x_n}$$

$$= \sum_{n=1}^N E_{z_n|x_n} [\log p(z_n=i) \cdot p(x_n | z_n=i) - \log p(z_n=i | x_n)]$$



General case:

$$\log E[X] \geq E[\log X]$$

# Posterior

$$\bullet p(z_n | x_n) = \frac{p(z_n, x_n)}{p(x_n)} = \frac{p(z_n) \cdot p(x_n | z_n)}{\sum_{j=1}^K p(z_n=j) \cdot p(x_n | z_n=j)}$$

$$\begin{aligned}\hat{z}_{n,i} &= p(z_n=i | x_n) = \frac{p(z_n=i) \cdot p(x_n | \mu_i, \Sigma_i)}{\sum_{j=1}^K p(z_n=j) \cdot p(x_n | \mu_j, \Sigma_j)} \\ &\approx \frac{w_i \cdot N(x_n, \mu_i, \Sigma_i)}{\sum_{j=1}^K w_j \cdot N(x_n, \mu_j, \Sigma_j)}\end{aligned}$$

# Update mixture weights

$$\begin{aligned}
 L(\theta) &= \sum_{n=1}^N \sum_{i=1}^K p(z_n=i | x_n) \cdot \log \frac{p(z_n=i) \cdot p(x_n | z_n=i)}{p(z_n=i | x_n)} \\
 &= \sum_{n=1}^N \sum_{i=1}^K \hat{z}_{n,i} \log \frac{w_i \cdot N(x_n, \mu_i, \Sigma_i)}{\hat{z}_{n,i}} \\
 &= \sum_{n=1}^N \sum_{i=1}^K \hat{z}_{n,i} (\log w_i + \log N(x_n, \mu_i, \Sigma_i) - \log \hat{z}_{n,i}) \\
 \text{s.t. } \sum_{j=1}^K w_j &= 1
 \end{aligned}$$

$\max f(x)$   
 $\text{s.t. } g(x) = 0$   
 Lagrangian  $\text{Lag}(x) = f(x) - \lambda g(x)$

$$\text{Lag}(\theta) = L(\theta) - \lambda \left( \sum_{j=1}^K w_j - 1 \right)$$

Optimality for  $w$ :

$$\frac{\partial \text{Lag}(\theta)}{\partial w_i} = 0$$

$$\hookrightarrow \sum_{n=1}^N \hat{z}_{n,i} \cdot \frac{1}{w_i} - \lambda = 0$$

$$\begin{aligned}
 w_i &= \frac{\sum_{n=1}^N \hat{z}_{n,i}}{\lambda} \\
 \Rightarrow \lambda &= \sum_{j=1}^K \sum_{n=1}^N \hat{z}_{n,j} \\
 w_i &= \frac{\sum_{n=1}^N \hat{z}_{n,i}}{\sum_{j=1}^K \sum_{n=1}^N \hat{z}_{n,j}}
 \end{aligned}$$

# Update mean and covariance

$$\log N(x_n, \mu_i, \Sigma_i) = -\frac{d}{2} \log 2\pi - \frac{1}{2} \log |\Sigma_i| - \frac{1}{2} (x_n - \mu_i)^T \Sigma_i^{-1} (x_n - \mu_i)$$

$$L(\theta) = \sum_{n=1}^N \sum_{i=1}^K \hat{z}_{n,i} \left[ -\frac{1}{2} \log |\Sigma_i| - \frac{1}{2} (x_n - \mu_i)^T \Sigma_i^{-1} (x_n - \mu_i) \right] + \dots$$

Optimality Condition

$$\frac{\partial L(\theta)}{\partial \mu_i} = \sum_{n=1}^N \cancel{\hat{z}_{n,i}} \cdot \left( -\frac{1}{2} \cdot 2 \cdot \Sigma_i^{-1} \cdot (x_n - \mu_i) \right) = 0$$

$$\mu_i = \frac{\sum_{n=1}^N \cancel{\hat{z}_{n,i}} \cdot x_n}{\sum_{n=1}^N \cancel{\hat{z}_{n,i}}}$$

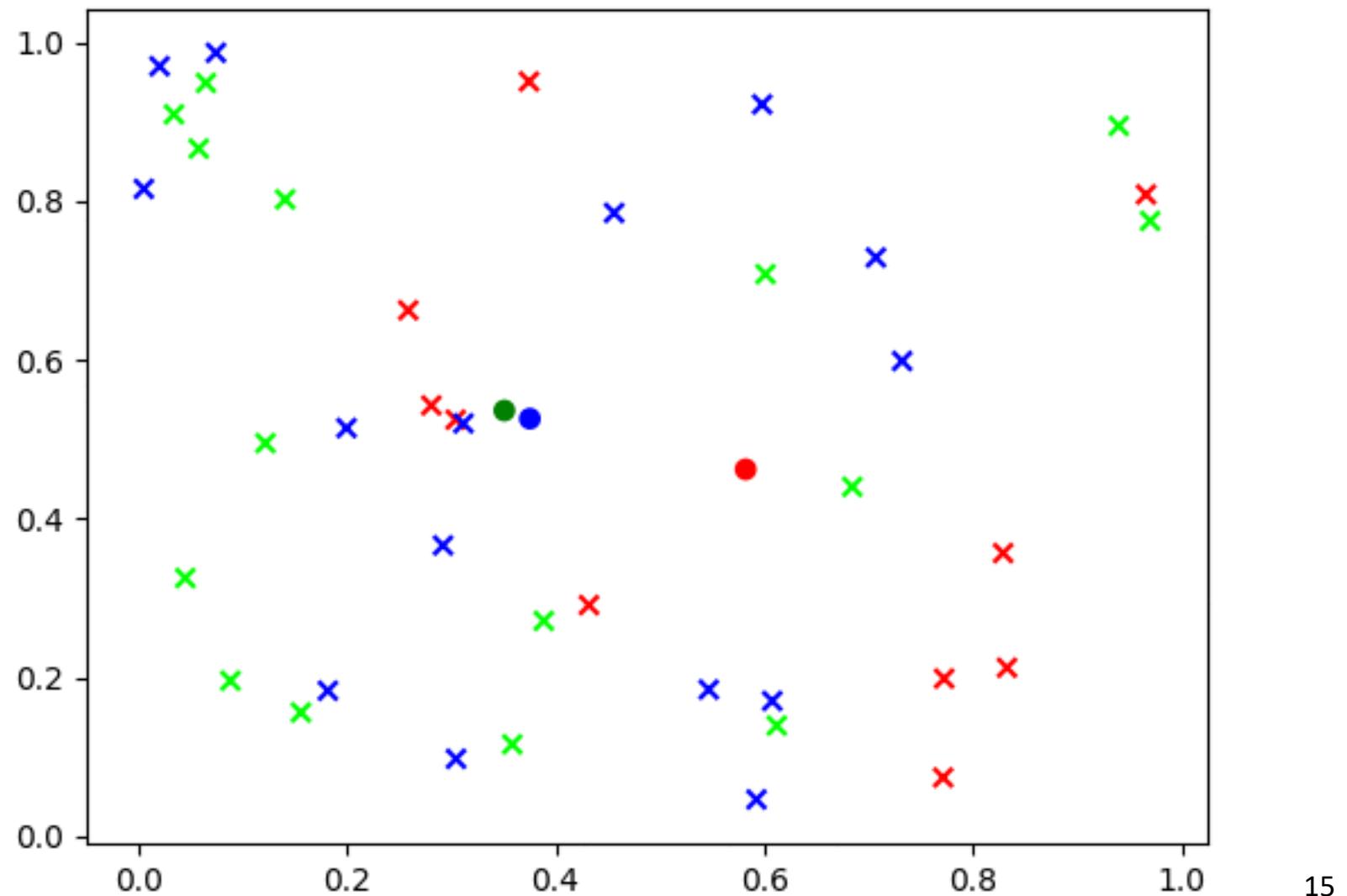
$$\frac{\partial L(\theta)}{\partial \Sigma_i} = -\frac{1}{2} \sum_{n=1}^N \hat{z}_{n,i} \left( \Sigma_i^{-1} - \Sigma_i^{-1} (x_n - \mu_i) \cdot (x_n - \mu_i)^T \Sigma_i^{-1} \right) = 0$$

$$\Sigma_i = \frac{\sum_{n=1}^N (x_n - \mu_i) (x_n - \mu_i)^T \cdot \hat{z}_{n,i}}{\sum_{n=1}^N \hat{z}_{n,i}}$$

# Summary of EM algorithm

- Observation:  $x_{1..N}$
- $\theta = \{w_{1..k}, \mu_{1..k}, \Sigma_{1..k}\}$
- Iterate until convergence
  1. E step: use  $X$  and current  $\theta$  to calculate  $p(z_{1..N} | x_{1..N}; \theta)$
  2. M step:
$$\theta \leftarrow \arg \max_{\theta} E_{p(z_{1..N} | x_{1..N}; \theta_{old})} \log p(x_n, z_n | \theta)$$
- Guaranteed to find local maximum
- Works for general mixture model

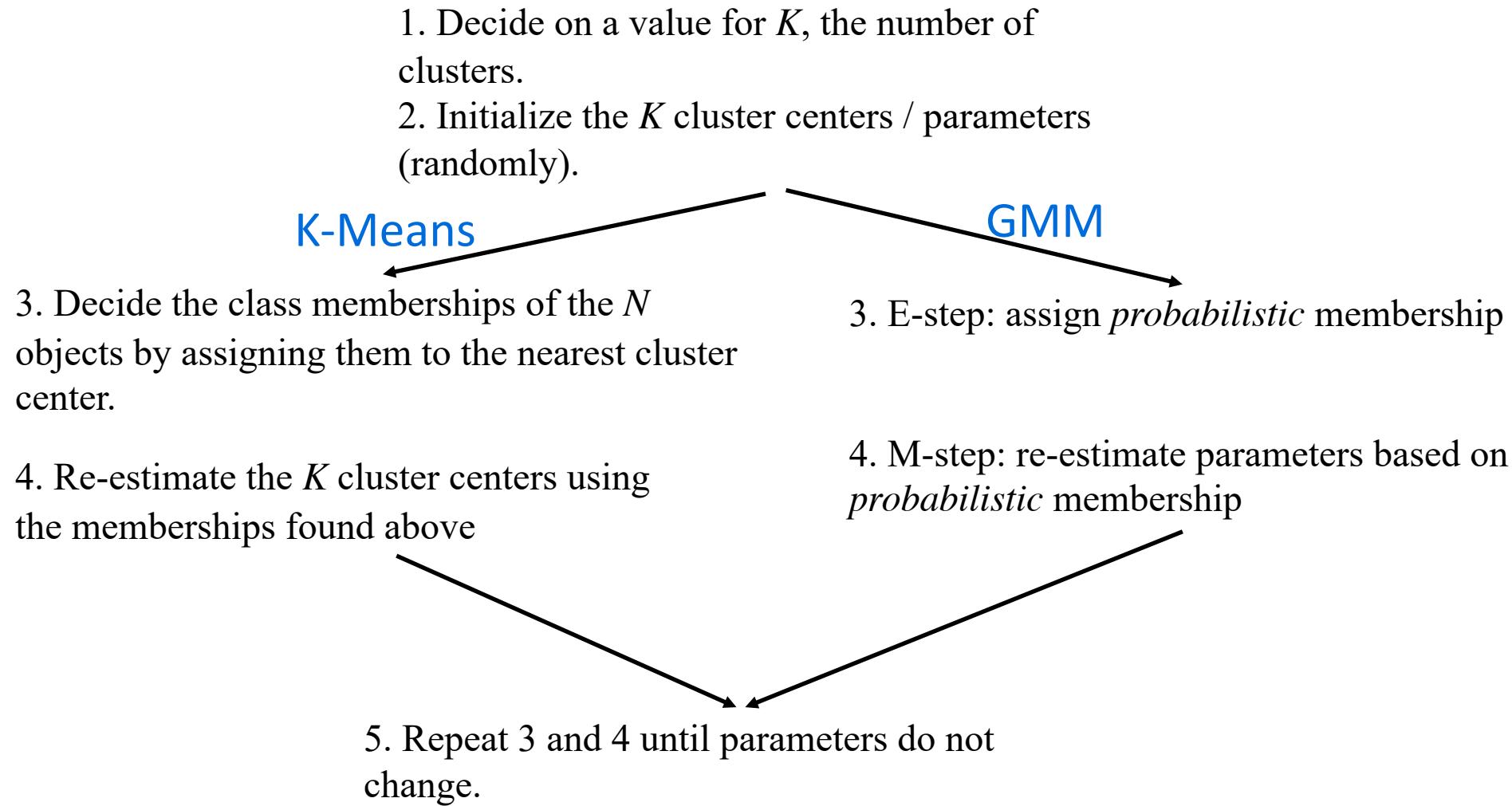
# Illustration of GMM



# Property of GMM

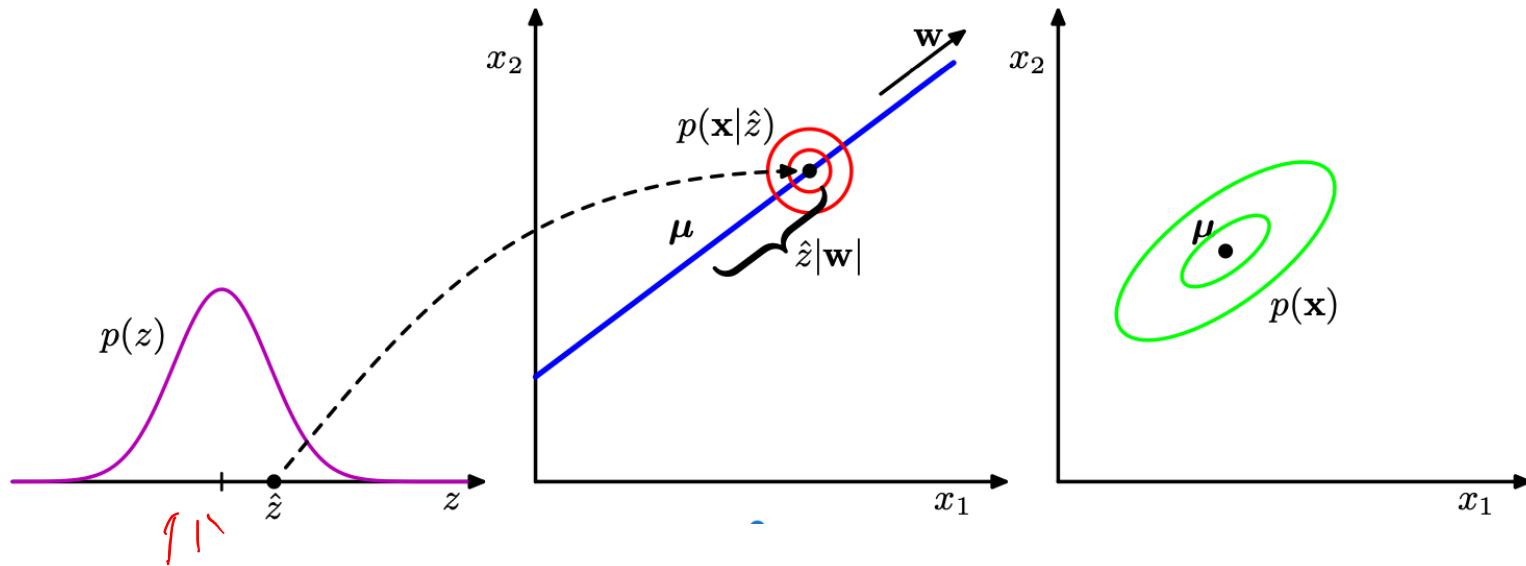
- Interpretable:
  - Participation weight of each data point from every component
- Generative:
  - Able to generate new data
- Handles missing values
- Efficient:  $O(TKN)$
- Local optimal:
  - Can be viewed as coordinate descent (why?)
- Need to specify K

# K-Means vs GMM



# Probabilistic PCA

- Continuous latent variable  $z \sim N(0, I)$
- Observation data  $x|z \sim N(W \cdot z + \mu, \sigma^2 I)$



# Learning Parameters for PPCA

- Again EM algorithm
- $\arg \max_{\theta} E_{p(z_{1..N}|x_{1..N};\theta_{old})} \log p(x_{1..N}, z_{1..N}|\theta)$

# A Variational View of EM

- $L(\theta) = \log p(X; \theta)$
- Introduce a variational distribution  $\underline{q(z; \phi)}$
- Variational bound for this data likelihood

$$= \log \int p(x, z; \theta) dz$$

$$= \log \int q(z; \phi) \cdot \frac{p(x, z; \theta)}{q(z; \phi)} dz$$

$$\xrightarrow{\text{Jensen's}} \int q(z; \phi) \cdot \underbrace{\log \frac{p(x, z; \theta)}{q(z; \phi)}}_{\text{(ELBO)}} dz \quad (\text{ELBO})$$

$$= \int q(z; \phi) \cdot \underbrace{\log \frac{p(x|z; \theta) \cdot p(z; \theta)}{q(z; \phi)}}_{\text{}} dz$$

$$= E_{q(z)} \log p(x|z; \theta) - KL(q(z; \phi) || p(z; \theta))$$

M-step:  
fix  $q, \phi$ ,  
estimate  $\theta$

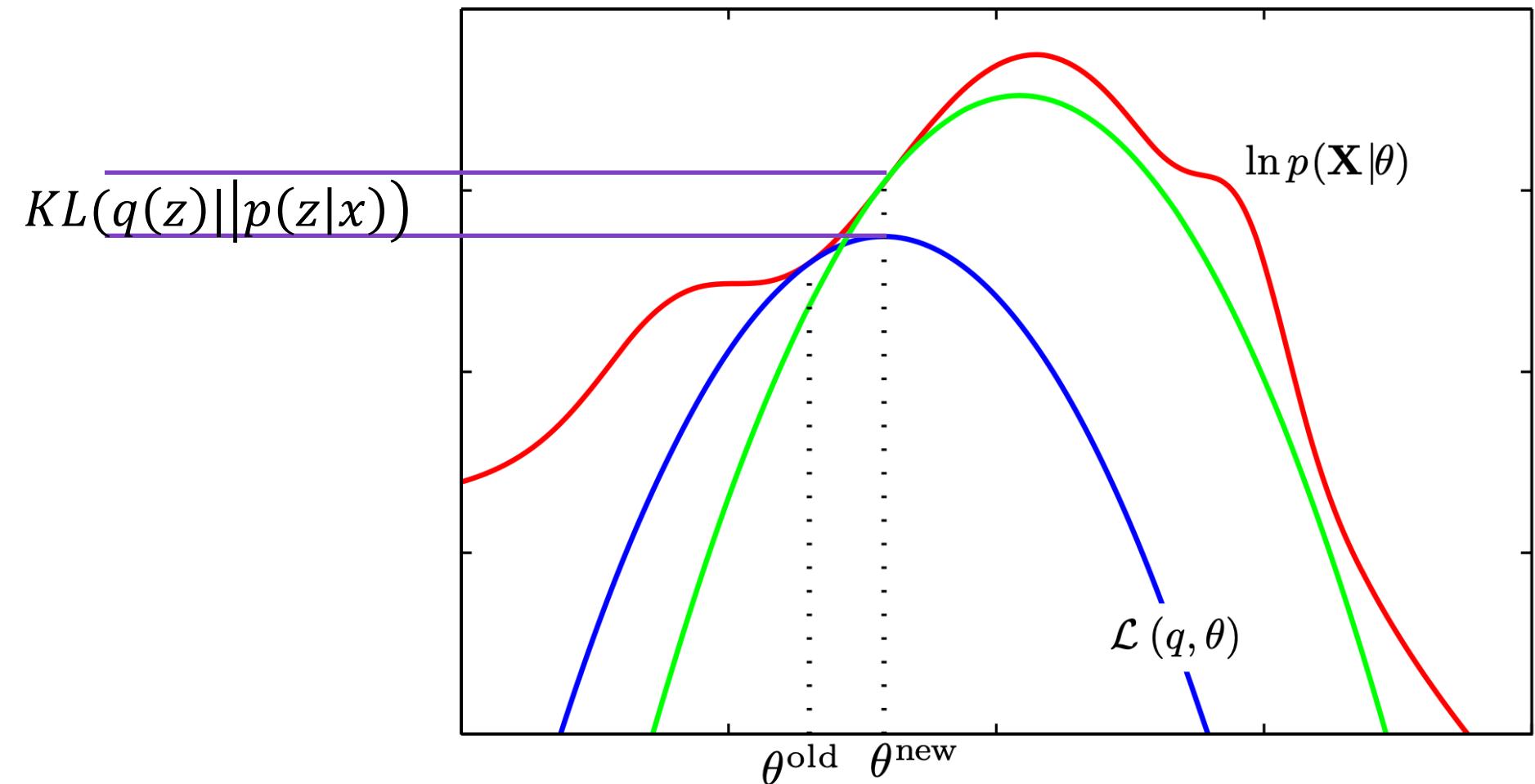
E-step:

$$\begin{aligned} &= \int q(z; \phi) \cdot \log \frac{p(z|x)}{q(z; \phi)} dz \\ &= -KL(q(z; \phi) || p(z|x; \theta)) \\ &\quad + \int q(z; \phi) \cdot \log p(x; \theta) dz \end{aligned}$$

$$\log p(x; \theta)$$

# What does EM actually do?

EM is coordinate-descent



# Summary

- Mixture Distribution: to build more complex distribution from simple ones
- Gaussian Mixture Model:  $k$  Gaussian components
- Expectation-Maximization: general for graphical models with latent variables
  - E-step: fix parameter, estimate posterior mean/variance
  - M-step: update parameter
- Probabilistic PCA: latent is continuous

# Recommended Reading

- PRML Chapter 9, 12.2

# Next up

- Dynamic Bayesian Network
- Linear Dynamical System